

Mark's Formula Sheet for Exam P

Discrete distributions

- Uniform, $U(m)$

- PMF: $f(x) = \frac{1}{m}$, for $x = 1, 2, \dots, m$

- $\mu = \frac{m+1}{2}$ and $\sigma^2 = \frac{m^2-1}{12}$

- Hypergeometric

- PMF: $f(x) = \frac{\binom{N_1}{x} \binom{N_2}{n-x}}{\binom{N}{n}}$

- x is the number of items from the sample of n items that are from group/type 1.

- $\mu = n \left(\frac{N_1}{N}\right)$ and $\sigma^2 = n \left(\frac{N_1}{N}\right) \left(\frac{N_2}{N}\right) \left(\frac{N-n}{N-1}\right)$

- Binomial, $b(n, p)$

- PMF: $f(x) = \binom{n}{x} p^x (1-p)^{n-x}$, for $x = 0, 1, \dots, n$

- x is the number of successes in n trials.

- $\mu = np$ and $\sigma^2 = np(1-p) = npq$

- MGF: $M(t) = [(1-p) + pe^t]^n = (q + pe^t)^n$

- Negative Binomial, $nb(r, p)$

- PMF: $f(x) = \binom{x-1}{r-1} p^r (1-p)^{x-r}$, for $x = r, r+1, r+2, \dots$

- x is the number of trials necessary to see r successes.

- $\mu = r \left(\frac{1}{p}\right) = \frac{r}{p}$ and $\sigma^2 = \frac{r(1-p)}{p^2} = \frac{rq}{p^2}$

- MGF: $M(t) = \frac{(pe^t)^r}{[1 - (1-p)e^t]^r} = \left(\frac{pe^t}{1 - qe^t}\right)^r$

- Geometric, $geo(p)$

- PMF: $f(x) = (1-p)^{x-1} p$, for $x = 1, 2, \dots$

- x is the number of trials necessary to see 1 success.

- CDF: $P(X \leq k) = 1 - (1-p)^k = 1 - q^k$ and $P(X > k) = (1-p)^k = q^k$

- $\mu = \frac{1}{p}$ and $\sigma^2 = \frac{1-p}{p^2} = \frac{q}{p^2}$

- MGF: $M(t) = \frac{pe^t}{1 - (1-p)e^t} = \frac{pe^t}{1 - qe^t}$

- Distribution is said to be “memoryless”, because $P(X > k + j | X > k) = P(X > j)$.

- Poisson

- PMF: $f(x) = \frac{\lambda^x e^{-\lambda}}{x!}$, for $x = 0, 1, 2, \dots$
- x is the number of changes in a unit of time or length.
- λ is the average number of changes in a unit of time or length in a Poisson process.
- CDF: $P(X \leq x) = e^{-\lambda}(1 + \lambda + \frac{\lambda^2}{2!} + \dots + \frac{\lambda^x}{x!})$
- $\mu = \sigma^2 = \lambda$
- MGF: $M(t) = e^{\lambda(e^t - 1)}$

Continuous Distributions

- Uniform, $U(a, b)$

- PDF: $f(x) = \frac{1}{b - a}$, for $a \leq x \leq b$
- CDF: $P(X \leq x) = \frac{x - a}{b - a}$, for $a \leq x \leq b$
- $\mu = \frac{a + b}{2}$ and $\sigma^2 = \frac{(b - a)^2}{12}$
- MGF: $M(t) = \frac{e^{tb} - e^{ta}}{t(b - a)}$, for $t \neq 0$, and $M(0) = 1$

- Exponential

- PDF: $f(x) = \frac{1}{\theta} e^{-x/\theta}$, for $x \geq 0$
- x is the waiting time we are experiencing to see one change occur.
- θ is the average waiting time between changes in a Poisson process. (Sometimes called the “hazard rate”.)
- CDF: $P(X \leq x) = 1 - e^{-x/\theta}$, for $x \geq 0$.
- $\mu = \theta$ and $\sigma^2 = \theta^2$
- MGF: $M(t) = \frac{1}{1 - \theta t}$
- Distribution is said to be “memoryless”, because $P(X \geq x_1 + x_2 | X \geq x_1) = P(X \geq x_2)$.

- Gamma

- PDF: $f(x) = \frac{1}{\Gamma(\alpha)\theta^\alpha} x^{\alpha-1} e^{-x/\theta} = \frac{1}{(\alpha - 1)! \theta^\alpha} x^{\alpha-1} e^{-x/\theta}$, for $x \geq 0$
- x is the waiting time we are experiencing to see α changes.
- θ is the average waiting time between changes in a Poisson process and α is the number of changes that we are waiting to see.
- $\mu = \alpha\theta$ and $\sigma^2 = \alpha\theta^2$
- MGF: $M(t) = \frac{1}{(1 - \theta t)^\alpha}$

- Chi-square (Gamma with $\theta = 2$ and $\alpha = \frac{r}{2}$)

- PDF: $f(x) = \frac{1}{\Gamma(r/2)2^{r/2}} x^{r/2-1} e^{-x/2}$, for $x \geq 0$
- $\mu = r$ and $\sigma^2 = 2r$
- MGF: $M(t) = \frac{1}{(1 - 2t)^{r/2}}$

- Normal, $N(\mu, \sigma^2)$

- PDF: $f(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-(x-\mu)^2/2\sigma^2}$

- MGF: $M(t) = e^{\mu t + \sigma^2 t^2/2}$

Integration formulas

- $\int p(x)e^{ax} dx = \frac{1}{a}p(x)e^{ax} - \frac{1}{a^2}p'(x)e^{ax} + \frac{1}{a^3}p''(x)e^{ax} - \dots$

- $\int_a^\infty x \left(\frac{1}{\theta}e^{-x/\theta}\right) dx = (a + \theta)e^{-a/\theta}$

- $\int_a^\infty x^2 \left(\frac{1}{\theta}e^{-x/\theta}\right) dx = ((a + \theta)^2 + \theta^2)e^{-a/\theta}$

Other Useful Facts

- $\sigma^2 = E[(X - \mu)^2] = E[X^2] - \mu^2 = M''(0) - M'(0)^2$

- $\text{Cov}(X, Y) = E[(X - \mu_x)(Y - \mu_y)] = E[XY] - \mu_x\mu_y$

- $\text{Cov}(X, Y) = \sigma_{xy} = \rho\sigma_x\sigma_y$

and

$$\rho = \frac{\sigma_{xy}}{\sigma_x\sigma_y}$$

- Least squares regression line: $y = \mu_y + \rho\frac{\sigma_y}{\sigma_x}(x - \mu_x)$

- When variables X_1, X_2, \dots, X_n are not pairwise independent, then

$$\text{Var}\left(\sum_{i=1}^n X_i\right) = \sum_{i=1}^n \sigma_i^2 + 2 \sum_{i<j} \sigma_{ij}$$

and

$$\text{Var}\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i^2 \sigma_i^2 + 2 \sum_{i<j} a_i a_j \sigma_{ij}$$

where σ_{ij} is the covariance of X_i and X_j .

- When X depends upon Y , $E[X] = E[E[X|Y]]$.
- When X depends upon Y , $\text{Var}(X) = E[\text{Var}(X|Y)] + \text{Var}(E[X|Y])$. (Called the “Total Variance” of X .)
- Chebyshev’s Inequality: For a random variable X having any distribution with finite mean μ and variance σ^2 , $P(|X - \mu| \geq k\sigma) \leq \frac{1}{k^2}$.

- For the variables X and Y having the joint PMF/PDF $f(x, y)$, the moment generating function for this distribution is

$$M(t_1, t_2) = E[e^{t_1 X + t_2 Y}] = E[e^{t_1 X} e^{t_2 Y}] = \sum_x \sum_y e^{t_1 x} e^{t_2 y} f(x, y)$$

- $\mu_x = M_{t_1}(0, 0)$ and $\mu_y = M_{t_2}(0, 0)$ (These are the first partial derivatives.)

- $E[X^2] = M_{t_1 t_1}(0, 0)$ and $E[Y^2] = M_{t_2 t_2}(0, 0)$ (These are the “pure” second partial derivatives.)

- $E[XY] = M_{t_1 t_2}(0, 0) = M_{t_2 t_1}(0, 0)$ (These are the “mixed” second partial derivatives.)

- Central Limit Theorem: As the sample size n grows,

- the distribution of $\sum_{i=1}^n X_i$ becomes approximately normal with mean $n\mu$ and variance $n\sigma^2$
- the distribution of $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$ becomes approximately normal with mean μ and variance $\frac{\sigma^2}{n}$.

• If X and Y are joint distributed with PMF $f(x, y)$, then

- the marginal distribution of X is given by $f_x(x) = \sum_y f(x, y)$

- the marginal distribution of Y is given by $f_y(y) = \sum_x f(x, y)$

- $f(x|y = y_0) = \frac{f(x, y_0)}{f_y(y_0)}$.

- $E[X|Y = y_0] = \sum_x x f(x|y = y_0) = \frac{\sum_x x f(x, y_0)}{f_y(y_0)}$